

Some generalized expectations associated with probability density functions of various multivariable distributions II

F.Y. AYANT¹

¹ Teacher in High School , France

ABSTRACT

In this paper, we employ the multivariable I-function defined by Prasad [3] to derive various generalized expectations associated with probability density functions, of various multivariate distributions (beta, gamma, Dirichlet distributions).

KEYWORDS : I-function of several variables, generalized expectations, probability density functions, multivariate distributions.

2010 Mathematics Subject Classification. 33C60, 82C31

1. Introduction

Recently Chandel and Gupta [2] are employed the multivariable H-function defined by Srivastava and Panda [4] in deriving expectations associated with probability density functions of various multivariate distributions. Here in the present document, we extend the work with the multivariable I-function defined by Prasad [3]. The I-function of several variables generalize the multivariable H-function defined by Srivastava et al [4] , itself is a generalisation of G-function of several variables. The multivariable I-function is defined in term of multiple Mellin-Barnes type integral :

$$I(z_1, z_2, \dots, z_r) = I_{p_2, q_2; p_3, q_3; \dots; p_r, q_r; p', q'; \dots; p^{(r)}, q^{(r)}}^{0, n_2; 0, n_3; \dots; 0, n_r; m', n'; \dots; m^{(r)}, n^{(r)}} \left(\begin{matrix} z_1 \\ \cdot \\ \cdot \\ \cdot \\ z_r \end{matrix} \middle| \begin{matrix} (a_{2j}; \alpha'_{2j}, \alpha''_{2j})_{1, p_2}; \dots; \\ \\ \\ (b_{2j}; \beta'_{2j}, \beta''_{2j})_{1, q_2}; \dots; \end{matrix} \right)$$

$$\left(\begin{matrix} (a_{rj}; \alpha'_{rj}, \dots, \alpha_{rj}^{(r)})_{1, p_r}; (a'_j, \alpha'_j)_{1, p'}; \dots; (a_j^{(r)}, \alpha_j^{(r)})_{1, p^{(r)}} \\ \\ (b_{rj}; \beta'_{rj}, \dots, \beta_{rj}^{(r)})_{1, q_r}; (b'_j, \beta'_j)_{1, q'}; \dots; (b_j^{(r)}, \beta_j^{(r)})_{1, q^{(r)}} \end{matrix} \right) \tag{1.1}$$

$$= \frac{1}{(2\pi\omega)^r} \int_{L_1} \dots \int_{L_r} \xi(t_1, \dots, t_r) \prod_{i=1}^s \phi_i(s_i) z_i^{s_i} ds_1 \dots ds_r \tag{1.2}$$

The defined integral of the above function, the existence and convergence conditions, see Y,N Prasad [3]. Throughout the present document, we assume that the existence and convergence conditions of the multivariable I-function.

The condition for absolute convergence of multiple Mellin-Barnes type contour (1.9) can be obtained by extension of the corresponding conditions for multivariable H-function given by as :

$$|arg z_k| < \frac{1}{2} \Omega_i^{(k)} \pi, \text{ where}$$

$$\Omega_i^{(k)} = \sum_{k=1}^{n^{(i)}} \alpha_k^{(i)} - \sum_{k=n^{(i)}+1}^{p^{(i)}} \alpha_k^{(i)} + \sum_{k=1}^{m^{(i)}} \beta_k^{(i)} - \sum_{k=m^{(i)}+1}^{q^{(i)}} \beta_k^{(i)} + \left(\sum_{k=1}^{n_2} \alpha_{2k}^{(i)} - \sum_{k=n_2+1}^{p_2} \alpha_{2k}^{(i)} \right) + \left(\sum_{k=1}^{n_r} \alpha_{rk}^{(i)} - \sum_{k=n_r+1}^{p_r} \alpha_{rk}^{(i)} \right) - \left(\sum_{k=1}^{q_2} \beta_{2k}^{(i)} + \sum_{k=1}^{q_3} \beta_{3k}^{(i)} + \dots + \sum_{k=1}^{q_r} \beta_{rk}^{(i)} \right) \quad (1.3)$$

where $i = 1, \dots, r$

The complex numbers z_i are not zero. Throughout this document, we assume the existence and absolute convergence conditions of the multivariable I-function.

We may establish the asymptotic expansion in the following convenient form :

$$I(z_1, \dots, z_r) = O(|z_1|^{\alpha'_1}, \dots, |z_r|^{\alpha'_r}), \max(|z_1|, \dots, |z_r|) \rightarrow 0$$

$$I(z_1, \dots, z_r) = O(|z_1|^{\beta'_1}, \dots, |z_r|^{\beta'_r}), \min(|z_1|, \dots, |z_r|) \rightarrow \infty$$

where $k = 1, \dots, r : \alpha'_k = \min[Re(b_j^{(k)} / \beta_j^{(k)})], j = 1, \dots, m_k$ and

$$\beta'_k = \max[Re((a_j^{(k)} - 1) / \alpha_j^{(k)})], j = 1, \dots, n_k$$

We will use these following notations in this paper :

$$U = p_2, q_2; p_3, q_3; \dots; p_{r-1}, q_{r-1}; V = 0, n_2; 0, n_3; \dots; 0, n_{s-1} \quad (1.4)$$

$$W = (p', q'); \dots; (p^{(r)}, q^{(r)}); X = (m', n'); \dots; (m^{(r)}, n^{(r)}) \quad (1.5)$$

$$A = (a_{2k}, \alpha'_{2k}, \alpha''_{2k}); \dots; (a_{(r-1)k}, \alpha'_{(r-1)k}, \alpha''_{(r-1)k}, \dots, \alpha^{(r-1)}_{(r-1)k}) \quad (1.6)$$

$$B = (b_{2k}, \beta'_{2k}, \beta''_{2k}); \dots; (b_{(r-1)k}, \beta'_{(r-1)k}, \beta''_{(r-1)k}, \dots, \beta^{(r-1)}_{(r-1)k}) \quad (1.7)$$

$$\mathfrak{A} = (a_{sk}; \alpha'_{rk}, \alpha''_{rk}, \dots, \alpha^r_{rk}) : \mathfrak{B} = (b_{rk}; \beta'_{rk}, \beta''_{rk}, \dots, \beta^r_{rk}) \quad (1.8)$$

$$A' = (a'_k, \alpha'_k)_{1,p'}; \dots; (a^{(r)}_k, \alpha^{(r)}_k)_{1,p^{(r)}}; B' = (b'_k, \beta'_k)_{1,p'}; \dots; (b^{(r)}_k, \beta^{(r)}_k)_{1,p^{(r)}} \quad (1.9)$$

The multivariable I-function write :

$$I(z_1, \dots, z_r) = I_{U;p_r,q_r;W}^{V;0,n_r;X} \left(\begin{matrix} z_1 & | & A; \mathfrak{A}; A' \\ \cdot & & \\ \cdot & & \\ \cdot & & \\ z_r & | & B; \mathfrak{B}; B' \end{matrix} \right) \quad (1.10)$$

2. Required results

In this section, we write following results

a) Liouville's theorem (see Chandel [1], page 83,eq .3.1)

$$\int_0^\infty \cdots \int_0^\infty f(x_1 + \cdots + x_r) x_1^{\mu_1-1} \cdots x_r^{\mu_r-1} dx_1 \cdots dx_r$$

$$= \frac{\Gamma(\mu_1) \cdots \Gamma(\mu_r)}{\Gamma(\mu_1 + \cdots + \mu_r)} \int_0^\infty f(t) t^{\mu_1 + \cdots + \mu_r - 1} dt, \text{Re}(\mu_i) > 0, i = 1, \dots, r \tag{2.1}$$

b)Gamma function

$$\Gamma(z) = \int_0^\infty e^{-t} t^{z-1} dt \quad \text{Re}(z) > 0 \tag{2.2}$$

c) Beta function

$$B(\alpha, \beta) = \int_0^\infty \frac{t^{\alpha-1}}{(1+t)^{\alpha+\beta}} dt, \text{Re}(\alpha) > 0, \text{Re}(\beta) > 0 \tag{2.3}$$

3. Multivariate Gamma distribution

$$f(x_1, \dots, x_r) = \frac{\Gamma(\mu_1) \cdots \Gamma(\mu_r) \lambda^{\mu + \mu_1 + \cdots + \mu_r}}{\Gamma(\mu + \mu_1 + \cdots + \mu_r)} e^{-(x_1 + \cdots + x_r)} (x_1 + \cdots + x_r)^\mu x_1^{\mu_1-1} \cdots x_r^{\mu_r-1} \tag{3.1}$$

where $\text{Re}(\lambda) > 0, \text{Re}(\mu) > 0, x_i \geq 0, \text{Re}(\mu_i) > 0, i = 1, \dots, r$ and $f(x_1, \dots, x_r) = 0$ elsewhere.

With the help of the equations (2.1) and (2.2), we obtain the following result.

$$\int_0^\infty \cdots \int_0^\infty f(x_1 + \cdots + x_r) dx_1 \cdots dx_r = 1 \tag{3.2}$$

Therefore $f(x_1, \dots, x_r)$ is a probability density function for multivariate gamma distribution.

4. Expectations associated with multivariate gamma distribution

Corresponding to density function $f(x_1, \dots, x_r)$ defined in (3.1), the expectation value of the function

$g(x_1, \dots, x_r)$ is defined as

$$\langle g(x_1, \dots, x_r) \rangle = \int_0^\infty \cdots \int_0^\infty f(x_1, \dots, x_r) g(x_1, \dots, x_r) dx_1 \cdots dx_r \tag{4.1}$$

Consider $g_1(x_1, \dots, x_r) = I_{U:p_r, q_r; X}^{V; 0, n_r; X} \left(\begin{matrix} z_1 x_1^{\alpha_1^1} \cdots x_r^{\alpha_r^1} (x_1 + \cdots + x_r)^{v_1} \\ \vdots \\ z_r x_1^{\alpha_1^r} \cdots x_r^{\alpha_r^r} (x_1 + \cdots + x_r)^{v_r} \end{matrix} \middle| \begin{matrix} A ; \mathfrak{A} ; A' \\ \vdots \\ B ; \mathfrak{B} ; B' \end{matrix} \right) \tag{4.2}$

where $I_{U:p_r, q_r; W}^{V; 0, n_r; X}$ is the I-function of r variables.

With the help of (4.1), (2.1) and (2.2), the expectation of $g_1(x_1, \dots, x_r)$ corresponding to density function $f(x_1, \dots, x_r)$ defined by (3.1), is given by

$$\begin{aligned} \langle g_1(x_1, \dots, x_r) \rangle &= \frac{\Gamma(\mu_1) \dots \Gamma(\mu_r) \lambda^{\mu + \mu_1 + \dots + \mu_r}}{\Gamma(\mu + \mu_1 + \dots + \mu_r)} \int_0^\infty \dots \int_0^\infty e^{-(x_1 + \dots + x_r)} (x_1 + \dots + x_r)^\mu \\ &\quad x_1^{\mu_1 - 1} \dots x_r^{\mu_r - 1} I_{U:p_r, q_r; W}^{V; 0, n_r; X} \left(\begin{matrix} z_1 x_1^{\alpha_1^1} \dots x_r^{\alpha_r^1} (x_1 + \dots + x_r)^{v_1} \\ \vdots \\ z_r x_1^{\alpha_1^r} \dots x_r^{\alpha_r^r} (x_1 + \dots + x_r)^{v_r} \end{matrix} \middle| \begin{matrix} A ; \mathfrak{A} ; A' \\ \vdots \\ B ; \mathfrak{B} ; B' \end{matrix} \right) dx_1 \dots dx_r \\ &= \frac{\Gamma(\mu_1) \dots \Gamma(\mu_r)}{\Gamma(\mu + \mu_1 + \dots + \mu_r)} I_{U:p_r+r+1, q_r+1; W}^{V; 0, n_r+r+1; X} \left(\begin{matrix} z_1 \\ \vdots \\ z_r \end{matrix} \middle| \begin{matrix} [1-\mu_j : \alpha_j^1, \dots, \alpha_j^r]_{1,r}, \\ \vdots \\ (1-\mu_1 - \dots - \mu_r; \alpha_1^1 + \dots + \alpha_r^1, \dots, \alpha_1^r + \dots + \alpha_r^r), \\ \\ (1 - \mu - \mu_1 - \dots - \mu_r; v_1 + \alpha_1^1 + \dots + \alpha_r^1, \dots, v_r + \alpha_1^r + \dots + \alpha_r^r), \mathfrak{A}; A' \\ \vdots \\ \mathfrak{B}; B' \end{matrix} \right) \end{aligned} \tag{4.3}$$

Provided that

- a) $Re(\lambda) > 0, Re(\mu_i) > 0, Re(v_i) > 0, Re(\alpha_j^i) > 0, x_i \geq 0, i, j = 1, \dots, r$
- b) $\left| \frac{arg z_k}{\lambda^{v_i + \alpha_1^i + \dots + \alpha_r^i}} \right| < \frac{1}{2} \Omega_i^{(k)} \pi$, where $\Omega_i^{(k)}$ is given in (1.3)
- c) $f(x_1, \dots, x_r) = 0$ elsewhere.

5. Multivariate Beta distribution

Consider the function

$$F(x_1, \dots, x_r) = \frac{\Gamma(\mu_1 + \dots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r) (x_1 + \dots + x_r)^\alpha}{\Gamma(\mu_1) \dots \Gamma(\mu_r) \Gamma(\beta) \Gamma(\alpha + \mu_1 + \dots + \mu_r)} \frac{1}{(1 + x_1 + \dots + x_r)^{\alpha + \beta + \mu_1 + \dots + \mu_r}} \tag{5.1}$$

where $Re(\alpha) > 0, Re(\beta) > 0, Re(\mu_i) > 0, i = 1, \dots, r$ and $F(x_1, \dots, x_r) = 0$ elsewhere. Now use to (2.1) and (2.3), we have

$$\int_0^\infty \dots \int_0^\infty F(x_1, \dots, x_r) dx_1 \dots dx_r = 1 \tag{5.2}$$

6. Expectation associated with multivariate Beta distribution

Corresponding to probability density function $F(x_1, \dots, x_r)$ defined by (5.1) consider the function

$$G(x_1, \dots, x_r) = G_1(x_1, \dots, x_r) = I_{U:p_r, q_r; W}^{V; 0, n_r; X} \left(\begin{array}{c} \frac{z_1 x_1^{\alpha_1^1} \dots x_r^{\alpha_r^1} (x_1 + \dots + x_r)^{\eta_1}}{(1 + x_1 + \dots + x_r)^{\alpha_1^1 + \dots + \alpha_r^1 + \eta_1}} \\ \vdots \\ \frac{z_r x_1^{\alpha_1^r} \dots x_r^{\alpha_r^r} (x_1 + \dots + x_r)^{\eta_r}}{(1 + x_1 + \dots + x_r)^{\alpha_1^r + \dots + \alpha_r^r + \eta_r}} \end{array} \middle| \begin{array}{l} A ; \mathfrak{A} ; A' \\ \vdots \\ B ; \mathfrak{B} ; B' \end{array} \right) \tag{6.1}$$

Therefore expectation of $G_1(x_1, \dots, x_r)$ is given by

$$\langle G_1(x_1, \dots, x_r) \rangle = \frac{\Gamma(\mu_1 + \dots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r) (x_1 + \dots + x_r)^\alpha}{\Gamma(\mu_1) \dots \Gamma(\mu_r) \Gamma(\beta) \Gamma(\alpha + \mu_1 + \dots + \mu_r)}$$

$$\int_0^\infty \dots \int_0^\infty \frac{(x_1 + \dots + x_r)^\alpha x_1^{\mu_1 - 1} \dots x_r^{\mu_r - 1}}{(1 + x_1 + \dots + x_r)^{\alpha + \beta + \mu_1 + \dots + \mu_r}} I_{U:p_r, q_r; W}^{V; 0, n_r; X} \left(\begin{array}{c} \frac{z_1 x_1^{\alpha_1^1} \dots x_r^{\alpha_r^1} (x_1 + \dots + x_r)^{\eta_1}}{(1 + x_1 + \dots + x_r)^{\alpha_1^1 + \dots + \alpha_r^1 + \eta_1}} \\ \vdots \\ \frac{z_r x_1^{\alpha_1^r} \dots x_r^{\alpha_r^r} (x_1 + \dots + x_r)^{\eta_r}}{(1 + x_1 + \dots + x_r)^{\alpha_1^r + \dots + \alpha_r^r + \eta_r}} \end{array} \middle| \begin{array}{l} A ; \mathfrak{A} ; A' \\ \vdots \\ B ; \mathfrak{B} ; B' \end{array} \right)$$

$$dx_1 \dots dx_r = \frac{\Gamma(\mu_1 + \dots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r) (x_1 + \dots + x_r)^\alpha}{\Gamma(\mu_1) \dots \Gamma(\mu_r) \Gamma(\alpha + \mu_1 + \dots + \mu_r)}$$

$$I_{U:p_r+r+1, q_r+2; W}^{V; 0, n_r+r+1; X} \left(\begin{array}{c} z_1 \\ \vdots \\ z_r \end{array} \middle| \begin{array}{l} A ; [1 - \mu_j : \alpha_j^1, \dots, \alpha_j^r]_{1, r} \\ \vdots \\ B ; (1 - \mu_1 - \dots - \mu_r ; \alpha_1^1 + \dots + \alpha_r^1, \dots, \alpha_1^r + \dots + \alpha_r^r) \end{array} \right)$$

$$\left. \begin{aligned} & (1 - \alpha - \mu_1 - \dots - \mu_r; \eta_1 + \alpha_1^1 + \dots + \alpha_r^1, \dots, \eta_r + \alpha_1^r + \dots + \alpha_r^r), \mathfrak{A}; A' \\ & \dots \\ & (1 - \beta - \alpha - \mu_1 - \dots - \mu_r; \eta_1 + \alpha_1^1 + \dots + \alpha_r^1, \dots, \eta_r + \alpha_1^r + \dots + \alpha_r^r), \mathfrak{B}; B' \end{aligned} \right\} \quad (6.2)$$

Provided that

a) $Re(\alpha) > 0, Re(\beta) > 0, Re(\mu_i) > 0, Re(v_i) > 0, Re(\alpha_i) > 0, x_i \geq 0, i = 1, \dots, r$

b) $|argz_k| < \frac{1}{2} \Omega_i^{(k)} \pi$, where $\Omega_i^{(k)}$ is given in (1.3)

c) $F(x_1, \dots, x_r) = 0$ elsewhere, $\alpha_i^j, \eta_i (i, j = 1, \dots, r)$ are non-negative real numbers

8. Multivariate Dirichlet distribution

Consider the function defined by

$$H(x_1, \dots, x_r) = \frac{\Gamma(\mu_1 + \dots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r) (x_1 + \dots + x_r)^\alpha}{\Gamma(\mu_1) \dots \Gamma(\mu_r) \Gamma(\beta) \Gamma(\alpha + \mu_1 + \dots + \mu_r)} x_1^{\mu_1-1} \dots x_r^{\mu_r-1} (x_1 + \dots + x_r)^\alpha [1 - (x_1 + \dots + x_r)]^{\beta-1} \quad (8.1)$$

at any point in the domain $x_i \geq 0, x_1 + \dots + x_r \leq 1, Re(\mu_i) > 0, Re(\alpha) > 0, Re(\beta) > 0, i = 1, \dots, r$ and $H(x_1, \dots, x_r) = 0$ elsewhere.

Therefore $\int_0^\infty \dots \int_0^\infty H(x_1, \dots, x_r) dx_1 \dots dx_r = \frac{\Gamma(\mu_1 + \dots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r)}{\Gamma(\mu_1) \dots \Gamma(\mu_r) \Gamma(\beta) \Gamma(\alpha + \mu_1 + \dots + \mu_r)} \int_0^\infty \dots \int_0^\infty x_1^{\mu_1-1} \dots x_r^{\mu_r-1} (x_1 + \dots + x_r)^\alpha [1 - (x_1 + \dots + x_r)]^{\beta-1} dx_1 \dots dx_r$

$$= \frac{\Gamma(\alpha + \beta + \mu_1 + \dots + \mu_r)}{\Gamma(\alpha + \mu_1 + \dots + \mu_r) \Gamma(\beta)} \int_0^\infty t^{\alpha + \mu_1 + \dots + \mu_r} (1 - t)^{\beta-1} dt = 1 \quad (8.2)$$

Hence $H(x_1, \dots, x_r)$ is probability density function for multivariate Dirichlet distribution.

9. Expectation associated with multivariate Dirichlet distribution

Corresponding to density function $H(x_1, \dots, x_r)$ defined in (8.1), the expectation value of the function is defined as $h(x_1, \dots, x_r)$

$$\langle h(x_1, \dots, x_r) \rangle = \int_0^\infty \dots \int_0^\infty h(x_1, \dots, x_r) H(x_1, \dots, x_r) dx_1 \dots dx_r \quad (9.1)$$

Further consider $h_1(x_1, \dots, x_r)$ defined by

$$I_{U:p_r, q_r; W}^{V; 0, n_r; X} \left(\begin{matrix} z_1 x_1^{\alpha_1^1} \cdots x_r^{\alpha_r^1} (x_1 + \cdots + x_r)^{\beta_1} (1 - x_1 - \cdots - x_r)^{\gamma_1} \\ \vdots \\ z_r x_1^{\alpha_1^r} \cdots x_r^{\alpha_r^r} (x_1 + \cdots + x_r)^{\beta_r} (1 - x_1 - \cdots - x_r)^{\gamma_r} \end{matrix} \middle| \begin{matrix} A ; \mathfrak{A} ; A' \\ \vdots \\ B ; \mathfrak{B} ; B' \end{matrix} \right) \tag{9.2}$$

Thus corresponding to probability density function $H(x_1, \dots, x_r)$ defined by (8.1), the expectation value of $h_1(x_1, \dots, x_r)$ is given by

$$\langle h_1(x_1, \dots, x_r) \rangle = \frac{\Gamma(\mu_1 + \cdots + \mu_r) \Gamma(\alpha + \beta + \mu_1 + \cdots + \mu_r)}{\Gamma(\mu_1) \cdots \Gamma(\mu_r) \Gamma(\beta) \Gamma(\alpha + \mu_1 + \cdots + \mu_r)} I_{U:p_r+r+2, q_r+2; W}^{V; 0, n_r+r+2; X} \left(\begin{matrix} z_1 \\ \vdots \\ z_r \end{matrix} \middle| \begin{matrix} [1-\mu_j : \alpha_j^1, \dots, \alpha_j^r]_{1,r}, (1-\beta; \gamma_1, \dots, \gamma_r) \\ \vdots \\ (1-\mu_1 - \cdots - \mu_r; \alpha_1^1 + \cdots + \alpha_r^1, \dots, \alpha_1^r + \cdots + \alpha_r^r) \\ \\ (1-\alpha - \mu_1 - \cdots - \mu_r; \beta_1 + \alpha_1^1 + \cdots + \alpha_r^1, \dots, \beta_r + \alpha_1^r + \cdots + \alpha_r^r), \mathfrak{A}; A' \\ \vdots \\ (1-\beta - \alpha - \mu_1 - \cdots - \mu_r; \beta_1 + \alpha_1^1 + \cdots + \alpha_r^1, \dots, \beta_r + \alpha_1^r + \cdots + \alpha_r^r), B ; B' \end{matrix} \right) \tag{9.3}$$

Provided that

- a) $Re(\alpha) > 0, Re(\beta) > 0, Re(\mu_i) > 0, x_i \geq 0, i = 1, \dots, r$
- b) $|arg z_k| < \frac{1}{2} \Omega_i^{(k)} \pi$, where $\Omega_i^{(k)}$ is given in (1.3)
- c) $F(x_1, \dots, x_r) = 0$ elsewhere, $\alpha_i^j, \gamma_i (i, j = 1, \dots, r)$ are non-negative real numbers

Remark

If $U = V = A = B = 0$, the multivariable I-function defined by Prasad degenerates in multivariable H-function defined by Srivastava et al [4], for more details, see Chandel and Gupta [2].

10. Conclusion

In this paper we have evaluated the expectation concerning three various multivariate distributions involving the multivariable I-function defined by Prasad [3]. The formulas established in this paper is of very general nature as it contains multivariable H-function, which is a general function of several variables studied so far. Thus, the integral established in this research work would serve as a key formula from which, upon specializing the parameters, as many as desired results involving the special functions of one and several variables can be obtained.

REFERENCES

- [1] Chandel R.C.S. The product of certain classical polynomials and the generalized Laplacian operator. *Ganita*, 20, (1969), page 79-87.
- [2] Chandel R.C.S and Gupta V. Some generalized expectations associated with probability density functions of various multivariate distributions. *Jnanabha* 42, (2012), page 125-142
- [3] Y.N. Prasad , Multivariable I-function , *Vijnana Parishad Anusandhan Patrika* 29 (1986) , page 231-237.
- [4] H.M. Srivastava and R.Panda. Some expansion theorems and generating relations for the H-function of several complex variables. *Comment. Math. Univ. St. Paul.* 24(1975), p.119-137

Personal adress : 411 Avenue Joseph Raynaud
Le parc Fleuri , Bat B
83140 , Six-Fours les plages
Tel : 06-83-12-49-68
Department : VAR
Country : FRANCE