# A Model for Optimal Reserve Inventory between Two Machines With Reference to the Distribution of Repair Time

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Abstract- In Inventory control theory, many suitable models for real life systems are constructed with the objective of determining the optimal inventory level. In a system where the machines are in series for producing the finished products. The reserve of semifinished products between two machines becomes unavoidable to minimize the idle time of machines in series. In this model the repair time of machines is assumed to be a random variable and it follows exponential distribution which satisfies the so-called change of distribution property. Also, the truncation point of the repair time distribution is itself a random variable and it follows mixed exponential distribution. Under this assumption an optimal reserve inventory is obtained.

**Keywords:** *Reserve Inventory, Change of distribution property, Repair time, Truncation point and Optimal reserve.* 

# I. INTRODUCTION

Determining the behavior of the inventory system in terms of the mathematical model brings out a determination of optimal inventory under different real-life circumstances. When working systems are considered, avoiding the breakdown of the system if quiet possible by keeping reserve inventory between the machines are in series. A system in which there are two machines M<sub>1</sub> and M<sub>2</sub> are in series. The output of the machine  $M_1$  is the input of the machine  $M_2$ . The breakdown of  $M_1$  causes the idle time of  $M_2$ , since there is no input to  $M_2$  from  $M_1$ . The idle time of M<sub>2</sub> is very costly and hence to avoid it, a reserve inventory is maintained in between  $M_1$  and  $M_2$ . If the reserve inventory is surplus in quantity, there is an inventory holding cost. If the reserve inventory slacks in its quantity, then it assumes the high idle time cost.

Hence, in order to balance these costs, an optimal reserve inventory must be maintained between these two machines. Therefore, the problem is to determine the optimal reserve between  $M_1$  and  $M_2$ . Here the repair time of  $M_1$  is considered as a random variable. The very basic model has been discussed by Hanssmann(1962)<sup>(1)</sup>. The extension over this model is discussed by Ramachandran and Sathyamoorthi (1981)<sup>(2)</sup>.

This type of models has been discussed by many authors under the assumptions that the repair time is a random variable. Sachithananthamet.al(2006)<sup>(3)</sup> discussed this model with the assumption that the probability function of the repair time undergoes a parametric change after the truncation point. The very basic concept of change of distribution property was discussed by Stangl(1995)<sup>(4)</sup>, Suresh Kumar(2006)<sup>(5)</sup> has used their concept in the stock model approach. The model for optimal reserve inventory between two machines under the assumption that the repair time of machine M<sub>1</sub> satisfies the change of property with distribution the truncation point being a random variable is derived by Sachithanantham et.al(2007)<sup>(6)</sup>. And the same model with a modification of the probability function of truncation point is discussed by Ramathilagam et. al(2014)<sup>(7)</sup>. The improvement over this model is being discussed in this paper, in which the truncation point is a random variable, which follows the exponential distribution with parameter  $\alpha$ .

# II. NOTATIONS

- h : Inventory holding cost/ unit/ unit-time.
- $d \quad : \quad Idle \ time \ cost \ due \ to \ M_2 \ / \ unit-time.$

- μ : Mean time interval between successive breakdown of machine M<sub>1</sub>, assuming exponential distribution of inter-arrival times.
- $t \quad : \quad Continuous \ random \ variable \ denoting \ the \\ repair \ time \ of \ M_1 \ with \ pdf \ g(.) \ and \ c.d.f \ G(.)$
- $r \quad : \quad Constant \ consumption \ rate \ of \ M_2 \ per \ unit \ of \ time.$
- S: Reserve inventory between  $M_1$  and  $M_2$ .
- T: Random variable denoting the idle time of  $M_2$
- $\theta$ : Parameter of exponential distribution before the truncation point.
- $\lambda$ : Parameter of exponential distribution after the truncation point.
- $\alpha$ : Parameter of exponential distribution.

# III. MODEL I

If T is a random variable denoting idle time of  $M_2$ , it can be seen that

$$T = \begin{cases} 0 & \text{if } \frac{s}{r} > t \\ t - \frac{s}{r} & \text{if } \frac{s}{r} \le t \end{cases}$$
(1)

The expected total cost of inventory holding and idle time of  $M_2$  per unit of time is given by

$$E(c) = hs + \frac{d}{\mu} \int_{\frac{s}{r}}^{\infty} \left( t - \frac{s}{r} \right) g(t) dt(2)$$

The optimal reserve  $\hat{S}$  can be obtained by solving the equation  $\frac{dE(c)}{ds} = 0.$ 

The expression for optimal reserve inventory is given by

$$G\left(\frac{\hat{S}}{r}\right) = 1 - \frac{r\mu h}{d}$$

This result is discussed in Hanssmann.F (1962). It may be observed that the above expression for optimal value of S has a constraint that  $\frac{h\mu r}{d}$  is always less than unity.Otherwise the solution is not a feasible one. Hence, a slight modification in the expression for the expected total cost can be incorporated as below.

# IV. MODEL II

In this model it is assumed that the repair time distribution satisfies the change of distribution property is basically discussed by Stangl(1995), suresh Kumar(2006). Under this assumption

$$g(t) = \begin{cases} g_1(t,\theta), & t \le X_0 \\ g_2(t,\lambda), & t > X_0 \end{cases}$$

where  $X_0$  is a random variable denoting the truncation point of repair time and it is distributed as exponential with parameter  $\alpha$ .

$$g(t,\theta) = \begin{cases} \theta e^{-\theta t} , & \text{if } t \leq X_0 \\ e^{-\theta x_0} e^{-(t-x_0)} \left( (t-x_0)^{\lambda-1} / \int_{1}^{t} \lambda \right) , & \text{if } t > X_0 \end{cases}$$

$$f(t) = g_1(t,\theta)P(t \le X_0) + g_2(t,\theta)P(t > X_0)$$
$$f(x_0) = \alpha e^{-\alpha x_0}, 0 \le x_0 < \infty$$

Here 
$$P(t \le X_0) = P(X_0 \ge t) = \int_t^\infty f(x_0) dx_0$$
  
$$= \int_t^\infty \alpha e^{-\alpha x_0} dx_0$$
$$= \alpha \left(\frac{e^{-\alpha x_0}}{-\alpha}\right)_t^\infty$$

$$\therefore P(t \le x_0) = e^{-\alpha t}$$

$$f(t) = g_1(t,\theta) e^{-\alpha t} + \int_0^t \alpha e^{-\alpha x_0} g_2(t,\theta) dx_0$$

It may be observed that the random variable 'T' defined in equation (1) also undergoes a parametric change and the average idle time of  $M_2$  is

$$\begin{split} E(T) &= \int_{\frac{s}{r}}^{\infty} (t - s/r) f(t) dt \\ &= \int_{\frac{s}{r}}^{\infty} (t - s/r) \left\{ g(t, \theta) e^{-\alpha t} + \int_{0}^{t} \alpha e^{-\alpha x_{0}} h(t, \theta) dx_{0} \right\} dt \\ &= \int_{\frac{s}{r}}^{\infty} (t - s/r) \left\{ \theta e^{-\theta t} e^{-\alpha t} dt \right\} + \alpha \int_{\frac{s}{r}}^{\infty} (t - s/r) e^{-\alpha x_{0}} e^{-\theta x_{0}} e^{-(t - x_{0})} \left( \frac{(t - x_{0})^{\lambda - 1}}{r} \right) dx_{0} dt \\ &= \theta \int_{\frac{s}{r}}^{\infty} (t - s/r) e^{-(\alpha + \theta)t} dt + \frac{\alpha}{r} \int_{\frac{s}{r}}^{\infty} \int_{0}^{t} (t - s/r) e^{-(\alpha + \theta)x_{0}} e^{-(t - x_{0})} (t - x_{0})^{\lambda - 1} dx_{0} dt \\ &= \theta I_{1} + \frac{\alpha}{r} \int_{1}^{\infty} (t - s/r) e^{-t} \int_{\frac{s}{r}}^{\infty} \left\{ \int_{0}^{t} e^{-(\alpha + \theta)x_{0}} e^{x_{0}} (t - x_{0})^{\lambda - 1} dx_{0} \right\} dt \\ &= \theta I_{1} + \frac{\alpha}{r} \int_{1}^{\infty} \int_{s/r}^{\infty} (t - s/r) e^{-t} \left[ \int_{0}^{t} e^{-(\alpha + \theta - 1)x_{0}} (t - x_{0})^{\lambda - 1} dx_{0} \right] dt \end{split}$$

$$\begin{aligned} \text{Let} \quad I_{\lambda} &= \frac{1}{\Gamma\lambda} \int_{0}^{t} e^{-(\alpha+\theta-1)x_{0}} (t-x_{0})^{\lambda-1} \, \mathrm{d}x_{0} \\ &= \frac{1}{\Gamma\lambda} \left\{ (t-x_{0})^{\lambda-1} \left[ \frac{e^{-(\alpha+\theta-1)x_{0}}}{-(\alpha+\theta-1)} \right]_{0}^{t} - \int_{0}^{t} \frac{e^{-(\alpha+\theta-1)x_{0}}}{-(\alpha+\theta-1)} (\lambda-1)(t-x_{0})^{\lambda-2} (-\mathrm{d}x_{0}) \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)} \int_{0}^{t} e^{-(\alpha+\theta-1)x_{0}} (t-x_{0})^{\lambda-2} \, \mathrm{d}x_{0} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)} \int_{0}^{t} e^{-(\alpha+\theta-1)x_{0}} (t-x_{0})^{\lambda-2} \, \mathrm{d}x_{0} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)} \left[ \frac{t^{\lambda-2}}{(\alpha+\theta-1)} - \frac{(\lambda-2)}{(\alpha+\theta-1)} I_{\lambda-3} \right] \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{2}} I_{\lambda-3} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{2}} \left[ \frac{t^{\lambda-3}}{(\alpha+\theta-1)} - \frac{(\lambda-3)}{(\alpha+\theta-1)} I_{\lambda-4} \right] \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{2}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{3}} I_{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)}{(\alpha+\theta-1)^{3}} t^{\lambda-3} - \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} \right\} \\ &= \frac{1}{\Gamma\lambda} \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)} - \frac{(\lambda-1)}{(\alpha+\theta-1)^{2}} t^{\lambda-2} + \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha+\theta-1)^{4}} t^{\lambda-4} + \frac{(\lambda-1)(\lambda-2)(\lambda-3)}{(\alpha$$

$$= \left\{ \frac{t^{\lambda-1}}{(\alpha+\theta-1)(\lambda-1)!} - \frac{t^{\lambda-2}}{(\alpha+\theta-1)^{2}(\lambda-2)!} + \frac{t^{\lambda-3}}{(\alpha+\theta-1)^{3}(\lambda-3)!} - \dots + \frac{t^{3}}{(\alpha+\theta-1)^{\lambda-3}3!} - \frac{t^{2}}{(\alpha+\theta-1)^{\lambda-2}2!} + \frac{t^{1}}{(\alpha+\theta-1)^{\lambda-1}1!} - \frac{t^{0}}{(\alpha+\theta-1)^{\lambda}0!} \right\}$$
$$= \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}t^{\lambda-i}}{(\alpha+\theta-1)^{i}(\lambda-i)!}$$
$$E(T) = \theta I_{1} + \alpha \int_{\frac{s}{r}}^{\infty} (t-s/r) e^{-t} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}t^{\lambda-i}}{(\alpha+\theta-1)^{i}(\lambda-i)!} \right] dt$$
$$E(T) = \theta \int_{\frac{s}{r}}^{\infty} (t-s/r) e^{-(\alpha+\theta)t} dt + \alpha \int_{\frac{s}{r}}^{\infty} (t-s/r) e^{-t} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}t^{\lambda-i}}{(\alpha+\theta-1)^{i}(\lambda-i)!} \right] dt$$

Thus, the expected total cost is

$$\begin{split} E(c) &= hs + \frac{d}{\mu} \ E(T) \\ E(c) &= hs + \frac{d}{\mu} \Biggl\{ \theta \int_{\frac{s}{r}}^{\infty} (t - s_{/r}) \ e^{-(\alpha + \theta)t} dt + \alpha \int_{\frac{s}{r}}^{\infty} (t - s_{/r}) \ e^{-t} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1} t^{\lambda - i}}{(\alpha + \theta - 1)^{i} (\lambda - i)!} \right] dt \Biggr\} \\ E(c) &= hs + \frac{d}{\mu} \{ \theta T_1 + \alpha T_2 \} \\ \frac{dE(c)}{ds} &= 0 \quad \Rightarrow h + \frac{\theta d}{\mu} \frac{dT_1}{ds} + \frac{\alpha d}{\mu} \frac{dT_2}{ds} = 0 \\ \frac{dT_1}{ds} &= \frac{d}{ds} \int_{\frac{s}{r}}^{\infty} (t - s_{/r}) \ e^{-(\alpha + \theta)t} dt \\ &= \left( -\frac{1}{r} \right) \int_{\frac{s}{r}}^{\infty} e^{-(\alpha + \theta)t} dt \\ &= \left( -\frac{1}{r} \right) \left[ \frac{e^{-(\alpha + \theta)t}}{-(\alpha + \theta - 1)} \right]_{\frac{s}{r}}^{\infty} \end{split}$$

$$\begin{split} \frac{d\mathbf{T}_1}{ds} &= \left[ \frac{-e^{-(\alpha+\theta)s/r}}{r(\alpha+\theta)} \right] \\ \frac{d\mathbf{T}_2}{ds} &= \frac{d}{ds} \int_{\frac{\pi}{r}}^{\infty} (t - s/r) \, e^{-t} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1} t^{\lambda-i}}{(\alpha+\theta-1)^i(\lambda-i)!} \right] dt \\ &= \left( -\frac{1}{r} \right) \int_{\frac{\pi}{r}}^{\infty} e^{-t} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1} t^{\lambda-i}}{(\alpha+\theta-1)^i(\lambda-i)!} \right] dt \\ &= \left( -\frac{1}{r} \right) \sum_{i=1}^{\lambda} (-1)^{i+1} \int_{\frac{\pi}{r}}^{\infty} e^{-t} \frac{t^{\lambda-i}}{(\alpha+\theta-1)^i(\lambda-i)!} dt \\ &= \left( -\frac{1}{r} \right) \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha+\theta-1)^i(\lambda-i)!} \int_{\frac{\pi}{r}}^{\infty} t^{\lambda-i} e^{-t} dt \\ &= \left( -\frac{1}{r} \right) \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha+\theta-1)^i(\lambda-i)!} e^{-\frac{\pi}{r}} \sum_{j=0}^{\lambda-i} \frac{\left( \frac{s}{r} \right)^j}{j!} \\ &= \left( -\frac{e^{-s/r}}{r} \right) \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha+\theta-1)^i(\lambda-i)!} \sum_{j=0}^{\lambda-i} \frac{(s/r)^j}{j!} \\ \frac{dE(c)}{ds} &= 0 \Rightarrow h + \frac{\theta d}{\mu} \left[ \frac{-e^{-(\alpha+\theta)s/r}}{r(\alpha+\theta)} \right] + \frac{\alpha d}{\mu} \left[ \left( -\frac{e^{-s/r}}{r} \right) \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha+\theta-1)^i(\lambda-i)!} \sum_{i=1}^{\lambda-i} \frac{(s/r)^j}{(\alpha+\theta-1)^i(\lambda-i)!} \sum_{j=0}^{\lambda-i} \frac{(s/r)^j}{j!} \right] = 0 \end{split}$$

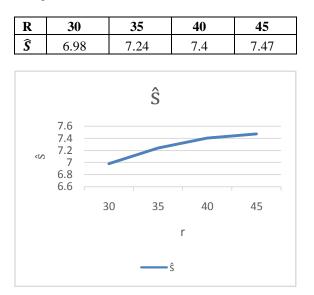
$$h = \frac{\theta d}{\mu r(\alpha + \theta)} e^{-\frac{(\alpha + \theta)s}{r}} + \frac{\alpha d}{\mu r} e^{-\frac{s}{r}} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha + \theta - 1)^{i}(\lambda - i)!} \sum_{j=0}^{\lambda - i} \frac{\left(\frac{s}{r}\right)^{j}}{j!} \right]$$
$$\frac{h\mu r}{d} = \frac{\theta}{(\alpha + \theta)} e^{-(\alpha + \theta)s/r} + \alpha e^{-s/r} \left[ \sum_{i=1}^{\lambda} \frac{(-1)^{i+1}}{(\alpha + \theta - 1)^{i}(\lambda - i)!} \sum_{j=0}^{\lambda - i} \frac{(s/r)^{j}}{j!} \right]$$

## V. NUMERICAL ILLUSTRATIONS

The variations in the values of optimal reserve inventory " $\hat{S}$  ", consequent on the changes in the parametersr, h, d,µ and  $\alpha$ have been studied by taking numerical illustrations. The tables and the corresponding graphs are given.

## Case (i)

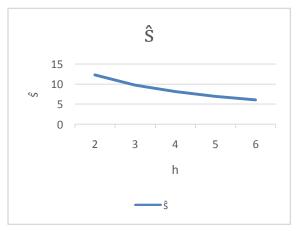
For h=5,d=3000, $\mu$ =2,  $\lambda$  =5, $\alpha$  =4,  $\theta$ =3, the optimal value of S is obtained and the variations in  $\hat{S}$  for the changes in the value of r.



### Case (ii)

For r=30, d=3000,  $\mu$ =2,  $\lambda$  =5, $\alpha$  =4, $\theta$ =3, the optimal value of S is obtained and the variations in  $\hat{S}$  for the changes in the value of h.

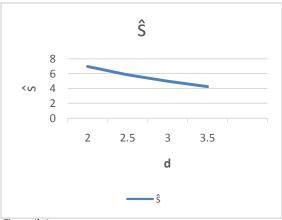
Η	2	3	4	5	6
Ŝ	12.34	9.75	8.14	6.98	6.06



#### Case (iii)

For r=30, h=5,  $\mu$ =2, $\lambda$  =5,  $\alpha$  =4,  $\theta$ =3, the optimal value of S is obtained and the variations in  $\hat{S}$  for the changes in the value of d.

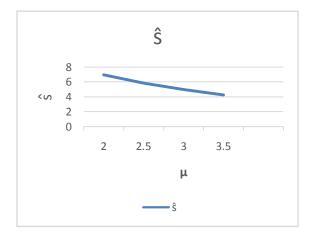
d	2500	3000	3500	4000
Ŝ	6.06	6.98	7.78	8.49



Case (iv)

For r=30, h=5, d=3000,  $\lambda = 5$ ,  $\alpha = 4$ ,  $\theta = 3$ , the optimal value of S is obtained and the variations in  $\hat{S}$  for the changes in the value of  $\mu$ .

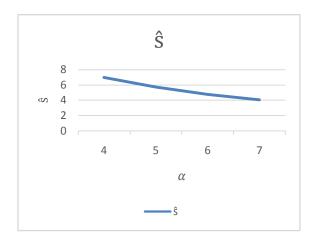
μ	2	2.5	3	3.5
Ŝ	6.98	5.86	4.98	4.25



#### Case (v)

For r=30, h=5, d=3000,  $\mu$ =2, $\lambda$  =5,  $\theta$ =3, the optimal value of S is obtained and the variations in  $\hat{S}$  for the changes in the value of  $\alpha$ .

α	4	5	6	7
Ŝ	6.98	5.71	4.76	4.04



#### VI. CONCLUSIONS

From the tables and graphs, it is observed that,

As r, the consumption rate of  $M_2$  increases, the optimal reserve inventory  $\hat{S}$  increases.

As h, the parameter of the holding cost increases, the optimal reserve inventory  $\hat{S}$  decreases. It is understood that, the model suggests small size inventory, when the inventory holding cost increases.

As d, the parameter of the idle time cost increases, the optimal reserve inventory  $\hat{S}$  decreases.

As  $\mu$ , the parameter of the Mean time interval between successive breakdown of machine  $M_1$  increases, the optimal reserve inventory  $\hat{S}$ decreases.

As $\alpha$ , the parameter of exponential distribution increases, the optimal reserve inventory  $\hat{S}$  decreases.

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