

Ratio Cum Median Based Modified Ratio Estimators with Known Skewness, and Coefficient of Variation

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Abstract

The present paper deals with some ratio cum median based modified ratio estimators forestimating the finite population mean with the linear combinations of known parameters of the auxiliary variable such as skewness, and Coefficient of variation. The efficiencies of the proposed estimators are assessed with that of simple random sampling without replacement (SRSWOR) sample mean and ratio estimator both by algebraically and numerically. From the numerical comparison it is observed that the proposed estimators perform better than the SRSWOR sample mean as well as the Ratio estimator.

Keywords: Auxiliary variable; Bias; Mean squared error; Natural population; Percentage relative efficiency; Simple random sampling.

1. INTRODUCTION

Let the finite population with N distinct and identifiable units be represented by $U = \{U_1, U_2, \dots, U_N\}$. Let Y and X denote the study and auxiliary variables and are taking the values Y_i and X_i respectively measured on the population unit $U_i, i = 1, 2, \dots, N$. In general the purpose of any sample survey is to estimate the unknown population mean $\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$ of the study variable by selecting a random sample of fixed size n from the population U with some desirable properties like unbiasedness, minimum variance etc. The widely used estimator is the simplerandom sampling without replacement (SRSWOR) sample mean which provides an unbiased estimator for the population mean in the absence of auxiliary variable. In the presence of an auxiliary variable one can always to improve the efficiency of SRSWOR sample mean by introducing Ratio estimator provided the auxiliary variable X is positively correlated with that of the study variable Y . For a more detailed discussion on ratio estimator and its related problems the readers are referred to Cochran (1977) and Murthy (1967). The efficiency of the ratio estimator can be improved further with the help of known parameters of the auxiliary variable such as, correlation coefficient, coefficient of variation, Skewness, Kurtosis, Quartiles etc. The resulting estimators are called in literature as modified ratio estimators. See for example Adepoju and Shittu (2013), Das and Tripathi (1978), Diana, Giordan and Perri (2011), Gupta and Shabbir (2008a,b), Kadilar and Cingi (2003a,b, 2004, 2005a,b, 2006a,b,c), Koyuncu (2012), Koyuncu and Kadilar (2009a,b), Shittu and Adepoju (2013), Singh and Agnihotri (2008), Singh and Tailor(2003), Sisodia and Dwivedi (1981), Subramani and Kumarapandiyam (2012a,b,c, 2013) and the references cited there in.

Recently Subramani (2013b) has introduced a new median based ratio estimator that uses the population median of the study variable Y and shown that the median based ratio estimator performs better than SRSWOR sample mean, ratio estimator and also the linear regression estimator. Later the median based ratio estimator of Subramani (2013b) has been extended and developed the median based modified ratio estimators by Subramani and Prabavathy (2014a,b,c, 2015). Recently Jayalakshmi et.al (2016) and Subramani et.al (2016) have introduced some ratio cum median based modified ratio estimators for estimation of finite population mean with known parameters of the auxiliary variable such as kurtosis, skewness, coefficient of variation and correlation coefficient and their linear combinations. In this paper an attempt is made to introduce some more ratio cum median based modified ratio estimators using some more linear combinations known parameters of the auxiliary variable such as skewness, coefficient of variation .

1.2 Notations to be used

- N – Population size
- n – Sample size

- $f = n/N$, Sampling fraction
- $\delta = \frac{1-f}{n}$, finite population correction
- $\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$, Population mean of the study variable
- $\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$, SRSWOR Sample mean of the study variable (Unbiased)
- $\bar{X} = \frac{1}{N} \sum_{i=1}^N X_i$, Population mean of the Auxiliary variable
- $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$, SRSWOR Sample mean of the Auxiliary variable (Unbiased)
- $S_y = \sqrt{\frac{1}{N-1} \sum_{i=1}^N (Y_i - \bar{Y})^2}$, Population standard deviation of Y
- $S_x = \sqrt{\frac{1}{N-1} \sum_{i=1}^N (X_i - \bar{X})^2}$, Population standard deviation of X
- $S_{xy} = E(X - \bar{X})(Y - \bar{Y}) = \frac{1}{N} \sum_{i=1}^N X_i Y_i - \bar{X}\bar{Y}$, Covariance between X and Y
- $C_x = \frac{S_x}{\bar{X}}$ & $C_y = \frac{S_y}{\bar{Y}}$ – Co-efficient of variations
- $\rho = \frac{S_{xy}}{S_x S_y}$ – Co-efficient of correlation between X and Y
- $\beta_{1(x)} = \frac{\mu_3}{\mu_2^2}$, Skewness of the auxiliary variable
- $\beta_{2(x)} = \frac{\mu_4}{\mu_2^2}$, Kurtosis of the auxiliary variable, where $\mu_r = \frac{1}{N} \sum_{i=1}^N (x_i - \bar{X})^r$
- Q_1 – First (lower) quartile of the auxiliary variable
- Q_3 – Third (upper) quartile of the auxiliary variable
- $Q_r = Q_3 - Q_1$, Inter-quartile range of the auxiliary variable
- $Q_d = \frac{Q_3 - Q_1}{2}$, Semi-quartile range of the auxiliary variable
- $Q_a = \frac{Q_3 + Q_1}{2}$ – Semi-quartile average of the auxiliary variable
- $B(\cdot)$ – Bias of the estimator
- $MSE(\cdot)$ – Mean squared error of the estimator
- $V(\cdot)$ – Variance of the estimator
- \bar{y} – SRSWOR Sample Mean (Unbiased)
- \hat{Y}_R – Ratio Estimator (Biased)
- \hat{Y}_i – i^{th} Existing modified ratio estimator of \bar{Y} (Biased)
- \hat{Y}_{pi} – i^{th} Proposed modified ratio estimator of \bar{Y} (Biased)

Existing Estimators

Simple Random Sampling without Replacement (SRSWOR)

In case of simple random sampling without replacement (SRSWOR), the sample mean \bar{y} is used to estimate population mean \bar{Y} which is an unbiased estimator and its variance is given below:

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i \quad (1)$$

$$V(\bar{y}) = \frac{(1-f)}{n} S_y^2 \quad (2)$$

where $f = \frac{n}{N}$, $S_y^2 = \frac{1}{N-1} \sum_{i=1}^N (Y_i - \bar{Y})^2$

Ratio Estimator for Estimation of Population Mean

In the presence of an auxiliary variable X and is positively correlated with the study variable Y the ratio estimator

$$\hat{Y}_R = \frac{\bar{y}}{\bar{x}} \bar{X} = \bar{R}\bar{X} \quad (3)$$

The mean squared error of \hat{Y}_R is given below:

$$MSE(\hat{Y}_R) = \bar{Y}^2 \{C'_{yy} + C'_{xx} - 2C'_{yx}\} \quad (4)$$

where $C'_{yy} = \frac{V(\bar{y})}{\bar{y}^2}$, $C'_{xx} = \frac{V(\bar{x})}{\bar{x}^2}$, $C'_{yx} = \frac{\text{Cov}(\bar{y}, \bar{x})}{\bar{X}\bar{Y}}$

2. PROPOSED ESTIMATORS

In this section, some more ratio cum median based modified ratio estimators with known linear combinations of the known parameters of the auxiliary variable likeskewness and coefficient of Variation, in line with the ratio cum median based modified ratio estimators by Jayalakshmi et al (2018). The proposed estimators together with their mean squared errors are given below:

Case i: The proposed estimator with known skewness β_1 and the coefficient of variation C_x is

$$\hat{Y}_{P_1} = \bar{y} \left\{ \alpha_1 \left(\frac{C_x M + \beta_1}{C_x m + \beta_1} \right) + \alpha_2 \left(\frac{C_x \bar{X} + \beta_1}{C_x \bar{x} + \beta_1} \right) \right\}$$

Case ii: The proposed estimator with known coefficient of variation C_x and the skewness β_1 is

$$\hat{Y}_{P_2} = \bar{y} \left\{ \alpha_1 \left(\frac{\beta_1 M + C_x}{\beta_1 m + C_x} \right) + \alpha_2 \left(\frac{\beta_1 \bar{X} + C_x}{\beta_1 \bar{x} + C_x} \right) \right\}$$

A. DERIVATION OF BIAS AND MEAN SQUARED ERRORS OF PROPOSED ESTIMATORS

The derivation of the bias and mean squared error of the proposed median based modified ratio estimator

\hat{Y}_{P_i} , $i = 1, 2$, are given below:

Consider $\hat{Y}_{P_i} = \bar{y} \left[\alpha_1 \left(\frac{M+T_i}{m+T_i} \right) + \alpha_2 \left(\frac{\bar{X}+T_i}{\bar{x}+T_i} \right) \right]$ (A1)

where $\alpha_1 + \alpha_2 = 1$

\bar{y} is the SRSWOR sample mean of the study variable Y

m is the sample median of the study variable Y

M is the population median of the study variable Y

T_i is the parameter or ratio of parameters of the auxiliary variable X

Let $e_0 = \frac{\bar{y}-Y}{\bar{Y}}$, $e_1 = \frac{m-M}{M}$ and $e_2 = \frac{\bar{x}-\bar{X}}{\bar{X}}$

$$\begin{aligned} \Rightarrow E(e_0) &= 0; E(e_1) = \frac{\bar{M}-M}{M} = \frac{B(m)}{M} \text{ and } E(e_2) = 0 \\ \Rightarrow E(e_0^2) &= \frac{V(\bar{y})}{\bar{Y}^2}; E(e_1^2) = \frac{V(m)}{M^2}; E(e_2^2) = \frac{V(\bar{x})}{\bar{X}^2} \\ E(e_0 e_1) &= \frac{\text{Cov}(\bar{y}, m)}{\bar{Y}M}; \end{aligned}$$

$$E(e_0 e_2) = \frac{\text{Cov}(\bar{y}, \bar{x})}{\bar{Y}\bar{X}}; E(e_1 e_2) = \frac{\text{Cov}(m, \bar{x})}{M\bar{X}}$$

The estimator \hat{Y}_{P_i} can be written in terms of e_0, e_1 and e_2 as

$$\begin{aligned} \hat{Y}_{P_i} &= \bar{Y}(1 + e_0) \left\{ \alpha_1 \left(\frac{M + T_i}{M(1 + e_1) + T_i} \right) + \alpha_2 \left(\frac{\bar{X} + T_i}{\bar{X}(1 + e_2) + T_i} \right) \right\} \\ \Rightarrow \hat{Y}_{P_i} &= \bar{Y}(1 + e_0) \left\{ \alpha_1 \left(\frac{M + T_i}{(M + T_i) + M e_1} \right) + \alpha_2 \left(\frac{\bar{X} + T_i}{(\bar{X} + T_i) + \bar{X} e_2} \right) \right\} \\ \Rightarrow \hat{Y}_{P_i} &= \bar{Y}(1 + e_0) \left\{ \alpha_1 \left(\frac{1}{1 + \left(\frac{M}{M+T_i} \right) e_1} \right) + \alpha_2 \left(\frac{1}{1 + \left(\frac{\bar{X}}{\bar{X}+T_i} \right) e_2} \right) \right\} \\ \Rightarrow \hat{Y}_{P_i} &= \bar{Y}(1 + e_0) \left\{ \alpha_1 \left(\frac{1}{1 + \theta_1 e_1} \right) + \alpha_2 \left(\frac{1}{1 + \phi_1 e_2} \right) \right\}; \end{aligned}$$

$$\text{where } \theta_i = \frac{M}{M + T_i}, \varphi_i = \frac{\bar{X}}{\bar{X} + T_i}$$

$$\Rightarrow \widehat{Y}_{p_i} = \bar{Y}(1 + e_0) \left\{ \alpha_1(1 + \theta_i e_1)^{-1} + \alpha_2(1 + \varphi_i e_2)^{-1} \right\}$$

Neglecting the terms of higher order, we have

$$\begin{aligned} \widehat{Y}_i &= \bar{Y} \{ \alpha_1(1 + e_0 - \theta_i e_1 - \theta_i e_0 e_1 + \theta_i^2 e_1^2) + \alpha_2(1 + e_0 - \varphi_i e_2 - \varphi_i e_0 e_2 + \varphi_i^2 e_2^2) \} \\ \Rightarrow \widehat{Y}_{p_i} &= \bar{Y} \{ 1 + e_0 + \alpha_1(-\theta_i e_1 - \theta_i e_0 e_1 + \theta_i^2 e_1^2) + \alpha_2(-\varphi_i e_2 - \varphi_i e_0 e_2 + \varphi_i^2 e_2^2) \} \\ \Rightarrow \widehat{Y}_{p_i} - \bar{Y} &= \bar{Y} \{ e_0 - \alpha_1(\theta_i e_1 + \theta_i e_0 e_1 - \theta_i^2 e_1^2) - \alpha_2(\varphi_i e_2 + \varphi_i e_0 e_2 - \varphi_i^2 e_2^2) \} \end{aligned} \tag{A2}$$

Taking expectations on both sides of (A2) one can have,

$$\begin{aligned} E(\widehat{Y}_{p_i} - \bar{Y}) &= \bar{Y} \{ E(e_0) - \alpha_1 E(\theta_i e_1 + \theta_i e_0 e_1 - \theta_i^2 e_1^2) - \alpha_2 E(\varphi_i e_2 + \varphi_i e_0 e_2 - \varphi_i^2 e_2^2) \} \\ \Rightarrow E(\widehat{Y}_{p_i} - \bar{Y}) &= \bar{Y} \left\{ \alpha_1 \left(\theta_i^2 \frac{V(m)}{M^2} - \theta_i \frac{B(m)}{M} - \theta_i \frac{\text{Cov}(\bar{y}, m)}{\bar{Y}M} \right) + \alpha_2 \left(\varphi_i^2 \frac{V(\bar{x})}{\bar{X}^2} - \varphi_i \frac{\text{Cov}(\bar{y}, \bar{x})}{\bar{Y}\bar{X}} \right) \right\} \\ \Rightarrow B(\widehat{Y}_{p_i}) &= \bar{Y} \left\{ \alpha_1 \left(\theta_i^2 C'_{mm} - \theta_i C'_{ym} - \theta_i \frac{B(m)}{M} \right) + \alpha_2 (\varphi_i^2 C'_{xx} - \varphi_i C'_{yx}) \right\} \end{aligned} \tag{A3}$$

Squaring on both sides of (A2), neglecting the terms of higher order and taking expectation on both sides one can get,

$$\begin{aligned} \text{MSE}(\widehat{Y}_{p_i}) &= E(\widehat{Y}_{p_i} - \bar{Y})^2 = \bar{Y}^2 E\{e_0 - \alpha_1 \theta_i e_1 - \alpha_2 \varphi_i e_2\}^2 \\ E(\widehat{Y}_{p_i} - \bar{Y})^2 &= \bar{Y}^2 \{ E(e_0^2) + \alpha_1^2 \theta_i^2 E(e_1^2) + \alpha_2^2 \varphi_i^2 E(e_2^2) - 2\alpha_1 \theta_i E(e_0 e_1) - 2\alpha_2 \varphi_i E(e_0 e_2) + 2\alpha_1 \alpha_2 \theta_i \varphi_i E(e_1 e_2) \} \end{aligned}$$

After a little algebra, the mean squared error of \widehat{Y}_{p_i} is obtained as

$$\begin{aligned} \text{MSE}(\widehat{Y}_{p_i}) &= \bar{Y}^2 \left\{ \frac{V(\bar{y})}{\bar{Y}^2} + \alpha_1^2 \theta_i^2 \frac{V(m)}{M^2} - 2\alpha_1 \theta_i \frac{\text{Cov}(\bar{y}, m)}{\bar{Y}M} + \alpha_2^2 \varphi_i^2 \frac{V(\bar{x})}{\bar{X}^2} - 2\alpha_2 \varphi_i \frac{\text{Cov}(\bar{y}, \bar{x})}{\bar{Y}\bar{X}} \right. \\ &\quad \left. + 2\alpha_1 \alpha_2 \theta_i \varphi_i \frac{\text{Cov}(m, \bar{x})}{M\bar{X}} \right\} \end{aligned}$$

That is,

$$\text{MSE}(\widehat{Y}_{p_i}) = \bar{Y}^2 \{ C'_{yy} + \alpha_1^2 \theta_i^2 C'_{mm} + \alpha_2^2 \varphi_i^2 C'_{xx} - 2\alpha_1 \theta_i C'_{ym} - 2\alpha_2 \varphi_i C'_{yx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \} \tag{A4}$$

Theorem 1: In SRSWOR, ratio cum median based modified ratio estimator

$\widehat{Y}_{P_5} = \bar{y} \left\{ \alpha_1 \left(\frac{C_x M + \beta_1}{C_x m + \beta_1} \right) + \alpha_2 \left(\frac{C_x \bar{X} + \beta_1}{C_x \bar{x} + \beta_1} \right) \right\}$ where $\alpha_1 + \alpha_2 = 1$ for the known parameter β_1 and C_x is not an unbiased estimator for its population mean \bar{Y} and its bias and MSE are respectively given as:

$$\begin{aligned} B(\widehat{Y}_{P_5}) &= \bar{Y} \left\{ \alpha_1 \left(\theta_i^2 C'_{mm} - \theta_i C'_{ym} - \theta_i \frac{B(m)}{M} \right) + \alpha_2 (\varphi_i^2 C'_{xx} - \varphi_i C'_{yx}) \right\} \\ \text{MSE}(\widehat{Y}_{P_5}) &= \bar{Y}^2 \{ C'_{yy} + \alpha_1^2 \theta_i^2 C'_{mm} + \alpha_2^2 \varphi_i^2 C'_{xx} - 2\alpha_1 \theta_i C'_{ym} - 2\alpha_2 \varphi_i C'_{yx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \}, \\ \text{where } \theta_i &= \frac{C_x M}{C_x M + \beta_1}, \varphi_i = \frac{C_x \bar{X}}{C_x \bar{X} + \beta_1} \end{aligned}$$

Proof: By replacing $T_i = \beta_1 / C_x$ in Theorem A the proof follows.

Theorem 2: In SRSWOR, ratio cum median based modified ratio estimator

$\widehat{Y}_{p_6} = \bar{y} \left\{ \alpha_1 \left(\frac{\beta_1 M + C_x}{\beta_1 m + C_x} \right) + \alpha_2 \left(\frac{\beta_1 \bar{X} + C_x}{\beta_1 \bar{x} + C_x} \right) \right\}$ where $\alpha_1 + \alpha_2 = 1$ for the known parameter C_x and β_1 is not an unbiased estimator for its population mean \bar{Y} and its bias and MSE are respectively given as:

$$B(\widehat{Y}_{p_6}) = \bar{Y} \left\{ \alpha_1 \left(\theta_i^2 C'_{mm} - \theta_i C'_{ym} - \theta_i \frac{B(m)}{M} \right) + \alpha_2 (\varphi_i^2 C'_{xx} - \varphi_i C'_{yx}) \right\}$$

$MSE(\widehat{Y}_{p_6}) = \bar{Y}^2 \{ C'_{yy} + \alpha_1^2 \theta_i^2 C'_{mm} + \alpha_2^2 \varphi_i^2 C'_{xx} - 2\alpha_1 \theta_i C'_{ym} - 2\alpha_2 \varphi_i C'_{yx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \}$, where

$$\theta_i = \frac{M}{M + \rho}, \varphi_i = \frac{\bar{X}}{\bar{X} + \rho}$$

Proof: By replacing $T_i = C_x / \beta_1$ Theorem A the proof follows.

The above proposed estimators are represented as a more generalized form as given below:

$$\widehat{Y}_{p_i} = \bar{y} \left\{ \alpha_1 \left(\frac{M + T_i}{m + T_i} \right) + \alpha_2 \left(\frac{\bar{X} + T_i}{\bar{x} + T_i} \right) \right\} \tag{5}$$

where $\alpha_1 + \alpha_2 = 1, \quad i = 1, 2$

The MSE of proposed estimator \widehat{Y}_{p_i} is given as

$$MSE(\widehat{Y}_{p_i}) = \bar{Y}^2 \{ C'_{yy} + \alpha_1^2 \theta_i^2 C'_{mm} + \alpha_2^2 \varphi_i^2 C'_{xx} - 2\alpha_1 \theta_i C'_{ym} - 2\alpha_2 \varphi_i C'_{yx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \} \tag{6}$$

$$\text{where } \theta_i = \frac{M}{M + T_i}, \varphi_i = \frac{\bar{X}}{\bar{X} + T_i}, C'_{xm} = \frac{\text{Cov}(\bar{x}, m)}{M\bar{X}}, T_1 = \beta_1 / C_x, T_2 = C_x / \beta_1$$

3.EFFICIENCY COMPARISON

The efficiencies of the proposed estimators are assessed with that of SRSWOR sample mean and ratio estimator in terms of variance/mean squared error. The results are as follows:

3.1 Comparison with that of SRSWOR sample mean

Comparing (6) and (2), it is noticed that the proposed estimators perform better than the SRSWOR sample mean That is

$$MSE(\widehat{Y}_{p_i}) \leq V(\bar{y}) \text{ if } \alpha_1^2 \theta_i^2 C'_{mm} + \alpha_2^2 \varphi_i^2 C'_{xx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \leq 2(\alpha_1 \theta_i C'_{ym} + \alpha_2 \varphi_i C'_{yx}) \tag{7}$$

3.2 Comparison with that of Ratio Estimator

Comparing (4.6) and (1.6), it is noticed that the proposed estimators perform better than the ratio estimator

That is $MSE(\widehat{Y}_{p_i}) \leq MSE(\widehat{Y}_R)$

$$\text{if } \alpha_1^2 \theta_i^2 C'_{mm} + (\alpha_2^2 \varphi_i^2 - 1) C'_{xx} + 2\alpha_1 \alpha_2 \theta_i \varphi_i C'_{xm} \leq 2[\alpha_1 \theta_i C'_{ym} + (\alpha_2 \varphi_i - 1) C'_{yx}] \tag{8}$$

3.3 NUMERICAL COMPARISON

The efficiencies of proposed ratio cum median based modified ratio estimators with that of existing estimators SRSWOR sample mean and ratio estimator are derived algebraically in the sections 3.1 and 3.2. To support it by means of numerical comparison, data of a natural population from Singh and Chaudhary (1986, page.177) has been considered.

Population Description

X= Area under Wheat in 1971 and Y= Area under Wheat in 1974

The population parameters computed for the above population is given below:

N= 34	n= 3	\bar{Y} = 856.4118
ρ = 0.4491	M= 767.5	\bar{X} = 208.8824
C_x = 0.7205	β_2 =2.9123	β_1 = 0.8732

The variance/mean squared error of the existing and proposed estimators at different values of α_1 and α_2 are given in the following table

Table 1: Mean Squared Errors for different values of α_1 and α_2

Existing Estimators			
SRSWOR Sample mean		\bar{y}	163356.41
Ratio Estimator		\hat{Y}_R	155579.71
Proposed Estimators			
α_1	α_2	\hat{Y}_{P_1}	\hat{Y}_{P_2}
0.1	0.9	140307.40	141938.20
0.2	0.8	129360.83	130052.11
0.3	0.7	119061.45	119650.02
0.4	0.6	110132.20	110744.34
0.5	0.5	101050.76	103216.29
0.6	0.4	97045.99	97339.09
0.7	0.3	92709.2	92844.20
0.8	0.2	89676.25	89690.97
0.9	0.1	88177.00	88326.88

From Table 1, it is observed that the proposed estimators have less mean squared errors than the SRSWOR sample mean and ratio estimator and hence one can easily conclude that the proposed ratio cum median based modified ratio estimators are performing better than the SRSWOR sample mean and ratio estimator for estimating the finite population mean of the study variable.

For the more clarity of assessing the performance, the percentage relative efficiencies (PRE) of the proposed estimators with respect to the existing estimators are obtained by using the formula

$PRE(e, p) = \frac{MSE(e)}{MSE(p)} * 100$ and are given in the following tables:

Table 2: PRE of proposed estimators with respect to SRSWOR sample mean

α_1	α_2	\widehat{Y}_{P_1}	\widehat{Y}_{P_2}
0.1	0.9	116.43	115.08
0.2	0.8	126.28	125.59
0.3	0.7	137.20	136.50
0.4	0.6	148.33	147.49
0.5	0.5	161.66	158.11
0.6	0.4	168.33	167.79
0.7	0.3	176.20	175.95
0.8	0.2	182.16	181.73
0.9	0.1	185.26	184.84

Table 3: PRE of proposed estimators with respect to Ratio Estimator

α_1	α_2	\widehat{Y}_{P_1}	\widehat{Y}_{P_2}
0.1	0.9	110.88	109.61
0.2	0.8	120.26	119.63
0.3	0.7	130.67	130.03
0.4	0.6	141.27	140.49
0.5	0.5	153.96	150.73
0.6	0.4	160.32	159.83
0.7	0.3	167.81	167.57
0.8	0.2	173.49	173.46
0.9	0.1	176.44	176.14

From Tables 2 and 3 it is observed that the PRE values of the proposed estimators with respect to SRSWOR sample mean and ratio estimator are greater than 100 and hence we conclude that the proposed estimators have greater efficiency.

- In fact the PREs are ranging from
 - 115.08 to 185.26 for the case of SRSWOR sample mean
 - 109.61 to 176.44 for the case of ratio estimator

CONCLUSION

The present paper deals with some more ratio cum median based modified ratio estimators with the linear combinations of the known parameters such as skewness and coefficient of variation of the auxiliary variable. The efficiencies of the proposed ratio cum median based modified ratio estimators are assessed algebraically as well as numerically with that of SRSWOR sample mean, and ratio estimator. Further it is shown from the numerical comparison that the PREs of proposed ratio cum median based modified ratio estimators with respect to the existing estimators are more than 100. Hence the proposed ratio cum median based modified ratio estimators may be recommended for the use of practical applications.

Acknowledgement

The Heading of the Acknowledgment section and the References section must not be numbered. Causal Productions wishes to acknowledge Michael Shell and the other contributors for developing and maintaining the IJMTT LaTeX style files which have been used in the preparation of this template. To see the list of contributors, Please refer to the top of the file IJMTT Tran.cls in the IJMTT LaTeX distribution

REFERENCES

- [1]. ADEPOJU, K.A. And SHITTU, O.L. (2013): On The Efficiency Of Ratio Estimator Based On Linear Combination Of Median, Coefficients Of Skewness And Kurtosis, American Journal Of Mathematics And Statistics, 3(3), 130-134
- [2]. COCHRAN, W. G. (1977): Sampling Techniques, Third Edition, Wiley Eastern Limited
- [3]. DAS, A.K. And TRIPATHI, T.P. (1978): Use Of Auxiliary Information In Estimating The Finite Population Variance, Sankhya, 40, 139-148
- [4]. DIANA, G., GIORDAN, M. And PERRI, P.F. (2011): An Improved Class Of Estimators For The Population Mean, Stat Methods Appl., 20, 123-140
- [5]. GUPTA, S. And SHABBIR, J. (2008): On Improvement In Estimating The Population Mean In Simple Random Sampling, Journal Of Applied Statistics, 35(5), 559–566
- [6]. JAYALAKSHMI, S.SUBRAMANI, J. And SRIJA, R. (2016): Ratio Cum Median Based Modified Ratio Estimators For The Estimation Of Finite Population Mean With Known Coefficient Of Variation And Correlation Coefficient, International Journal Of Computer And Mathematical Sciences, Intern. Jour. Computers And Math. Sci., 5(6), 122-127
- [7]. KADILAR, C. And CINGI, H. (2004): Ratio Estimators In Simple Random Sampling, Applied Mathematics And Computation, 151, 893-902
- [8]. KADILAR, C. And CINGI, H. (2006a): An Improvement In Estimating The Population Mean By Using The Correlation Coefficient, Hacettepe Journal Of Mathematics And Statistics, 35 (1), 103-109
- [9]. KADILAR, C. And CINGI, H. (2006b): Improvement In Estimating The Population Mean In Simple Random Sampling, Applied Mathematics Letters, 19, 75-79
- [10]. KOYUNCU, N. (2012): Efficient Estimators Of Population Mean Using Auxiliary Attributes, Applied Mathematics And Computation, 218, 10900-10905
- [11]. KOYUNCU, N. And KADILAR, C. (2009): Efficient Estimators For The Population Mean, Hacettepe Journal Of Mathematics And Statistics, 38(2), 217-225
- [12]. MURTHY, M.N. (1967): Sampling Theory And Methods, Statistical Publishing Society, Calcutta, Indiashittu, O.L. And ADEPOJU, K.A. (2013): On The Efficiency Of Some Modified Ratio And Product Estimators – The Optimal Approach, American Journal Of Mathematics And Statistics, 3(5), 296-299
- [13]. SINGH, D. AND CHAUDHARY, F.S (1986): Theory And Analysis Of Sample Survey Designs, New Age International Publisher
- [14]. SINGH, H.P. And AGNIHOTRI, N. (2008): A General Procedure Of Estimating Population Mean Using Auxiliary Information In Sample Surveys, Statistics In Transition, 9(1), 71–87
- [15]. SINGH, H.P. And TAILOR, R. (2003): Use Of Known Correlation Co-Efficient In Estimating The Finite Population Means, Statistics In Transition, Vol. 6 (4), 555-560
- [16]. SISODIA, B.V.S. And DWIVEDI, V.K. (1981): A Modified Ratio Estimator Using Co-Efficient Of Variation Of Auxiliary Variable, Journal Of The Indian Society Of Agricultural Statistics, Vol. 33(2), 13-18

- [17]. SISODIA, B.V.S. And DWIVEDI, V.K. (1981): A Modified Ratio Estimator Using Co-Efficient Of Variation Of Auxiliary Variable, *Journal Of The Indian Society Of Agricultural Statistics*, Vol. 33(2), 13-18

- [18]. SRIJA, R, SUBRAMANI, J. And JAYALAKSHMI, S. (2016): Ratio Cum Median Based Modified Ratio Estimators For The Estimation Of Finite Population Mean With Known Skewness, Kurtosis And Correlation Coefficient, *International Journal Of Computer And Mathematical Sciences*, Intern. Jour. Computers And Math. Sci., 5(7), 29-34

- [19]. SUBRAMANI, J. And KUMARAPANDIYAN, G. (2012a): Modified Ratio Estimators For Population Mean Using Function Of Quartiles Of Auxiliary Variable, *Bonfring International Journal Of Industrial Engineering And Management Science*, Vol. 2(2), 19-23

- [20]. SUBRAMANI, J. And KUMARAPANDIYAN, G. (2012b): Modified Ratio Estimators Using Known Median And Co-Efficient Of Kurtosis, *American Journal Of Mathematics And Statistics*, Vol. 2(4), 95-100

- [21]. SUBRAMANI, J. And PRABAVATHY, G. (2014a): Some Modified Ratio Estimators With Known Coefficient Of Variation And Correlation Coefficient, *Proceedings Of International Conference On Recent Developments In Statistical Theory And Practice (ICRDSTAP)*, 192-200, ISBN 978-93-83459-13-1

- [22]. SUBRAMANI, J. And PRABAVATHY, G. (2014b): Median Based Modified Ratio Estimators With Linear Combinations Of Population Mean And Median Of An Auxiliary Variable, *Journal Of Reliability And Statistical Studies*, Vol. 7(1), 01-10.

- [23]. SUBRAMANI, J. SRIJA, R. And JAYALAKSHMI, S. (2016): Ratio Cum Median Based Modified Ratio Estimators For The Estimation Of Finite Population Mean With Known Skewness And Kurtosis, *International Journal Of Computer And Mathematical Sciences*, Intern. Jour. Computers And Math. Sci., 5(6), 115-121

- [24]. JAYALAKSHMI, S. SUBRAMANI, J. And SRIJA, R. (2016): Ratio Cum Median Based Modified Ratio Estimators For The Estimation Of Finite Population Mean With Known Skewness, Kurtosis And Correlation Coefficient, *International Journal Of Computer And Mathematical Sciences*, Intern. Jour. Computers And Math. Sci., 5(6), 128-133

- [25]. SRIJA, R. SUBRAMANI, J (2018): Modified Ratio Cum Product Estimators For Finite Population Mean With Known Median And Mean, *International Journal Of Mathematics Trends And Technology(IJMTT)* Accepted For Publication