A FAMILY OF LOGARITHMIC ESTIMATORS FOR POPULATION VARIANCE UNDER DOUBLE SAMPLING

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Abstract

In this paper, an improved estimator for variance has been proposed to improvise the log-type estimators proposed by Kumari (2019). These classes of estimators provide a better alternative to the classes of estimators provided by Kumari (2019) as well as to some other commonly used estimators available in literature. A numerical study is included to support the use of the suggested classes of estimators.

1 Introduction

In sample surveys, auxiliary information always plays a vital role in better estimation of the parameter under investigation. This information can be used to improve the precision of the estimators. The suitable utilization of this auxiliary information can reduce the MSE of the sample mean, thus resulting in more efficient estimators. This paper, is an attempt to extend the powerful Searls approach to the traditional estimators using auxiliary information regarding to variables in simple random sampling. Many authors like, Singh et al. (1973), Das and Tripathi (1978), Sisodia and Dwivedi (1981), Isaki (1983), Bahl and Tuteja (1991), Prasad and Singh (1992), Swain (1994), Garcia and Cebrian (1996), Upadhaya and Singh (2001), Kalidar and Cingi (2006a, 2006b); Gupta and Shabbir (2006, 2007), Yadav and Kadilar (2013, 2014) had proposed an improved ratio estimators using Searls type estimators. Recently, Kumari et al. (2017, 2018) among others; Bhushan et al. (2015) have made the use of logarithmic relationship between the auxiliary variable and study variable as logarithm function which is very common in various branches of science as well as non-science disciplines.

In this paper, some improved logarithmic estimators are proposed for improving the efficiency of the Bhushan and Kumari (2019) estimators as these classes of estimators are expected to improve the mean squared error. The proposed estimators would work considerably well in case when the study variable is logarithmically related to the auxiliary variable.

Consider a finite population $U = U_1, U_2, ..., U_N$ of size N. From this population we draw a large sample of size n' to estimate the population parameters. Again from n' a sample of size n is drawn according to simple random sampling without replacement (SRSWOR) under double sampling. Let y_i and x_i denotes the values of the study and auxiliary variables for the *i*th unit (i = 1, 2, ..., N), of the population. Further, let $\bar{x'}$ be the mean of larger sample, \bar{y} and \bar{x} be the sample means and $s_x^{2'} = \sum_{i=1}^n (x_i - \bar{x'})^2/(n-1)$ be the larger sample variance and $s_y^2 = \sum_{i=1}^n (y_i - \bar{y})^2/(n-1)$ and $s_x^2 = \sum_{i=1}^n (x_i - \bar{x})^2/(n-1)$ be the sample variance of the study and auxiliary variables respectively.

2 The proposed estimators

We propose the following new classes of log-type estimators for the population variance S_u^2 as:

$$T_1^* = w_1 s_y^2 \left[1 + \log\left(\frac{s_x'^2}{s_x^2}\right) \right]^{a_1}$$
(2.1)

$$T_2^* = w_2 s_y^2 \left[1 + a_2 \log \left(\frac{s_x'^2}{s_x^2} \right) \right]$$
(2.2)

$$T_3^* = w_3 s_y^2 \left[1 + \log\left(\frac{s_x'^2}{s_x^{*2}}\right) \right]^{a_3}$$
(2.3)

$$T_4^* = w_4 s_y^2 \left[1 + a_4 \log\left(\frac{s_x'^2}{s_x^{*2}}\right) \right]$$
(2.4)

where

$$s_x^{*2} = as_x^2 + b$$

 $s_x^{'*2} = as_x^{'2} + b$

such that $a \neq 0$, b are either real numbers or functions of the known parameters of the auxiliary variables x such as the standard deviations S_x , coefficient of variation C_x , coefficient of kurtosis b_{2x} , coefficient of skewness b_{1x} and correlation coefficient ρ of the population. If $a_i = 0$, then the proposed estimator becomes the usual per unit variance estimator s_y^2 . If $a_i = +1$, then the proposed estimator works as ratio type estimator and if $a_i = -1$, then the proposed estimator type estimator having efficiency conditions equivalent to that of generalized product and ratio estimators respectively.

3 Properties of proposed estimators

In this paper, the mean square error(s) of all the estimators are considered up to the terms of order n^{-1} .

Theorem 1. MSE of T_1^* are given by

$$MSE(T_1^*) = S_y^4 w_1^2 \left[1 + Ib_{2y}^* + (I - I')(2a_1^2b_{2x}^* + 4a_1I_{22}^* - 2a_1b_{2x}^*) \right] - 2S_y^4 w_1 \left[1 + (I - I')(a_1I_{22}^* - a_1b_{2x}^* + \frac{a_1^2b_{2x}^*}{2}) \right] + S_y^4$$

Proof. Consider the estimator,

$$\begin{split} T_1^* &= w_1 s_y^2 \left[1 + \log \left(\frac{s_x'^2}{s_x^2} \right) \right]^{a_1} \\ &= w_1 S_y^2 \left(1 + \epsilon_0 \right) \left[1 + \log \left(1 + \epsilon_1' \right) \left(1 + \epsilon_1 \right)^{-1} \right]^{a_1} \\ &= w_1 S_y^2 \left[1 + \epsilon_0 + a_1 \left(\epsilon_1' - \epsilon_1 + \epsilon_1^2 - \epsilon_1 \epsilon_1' \right) - \frac{a_1 \left(\epsilon_1' - \epsilon_1 \right)^2}{2} + \frac{a_1 \left(a_1 - 1 \right)}{2} \left(\epsilon_1' - \epsilon_1 \right)^2 \right. \\ &+ a_1 \left(\epsilon_0 \epsilon_1' - \epsilon_0 \epsilon_1 \right) \right] \\ &= \left(w_1 - 1 \right) S_y^2 + w_1 S_y^2 \left[\epsilon_0 + a_1 \left(\epsilon_1' - \epsilon_1 + \epsilon_1^2 - \epsilon_1 \epsilon_1' \right) - \frac{a_1 \left(\epsilon_1' - \epsilon_1 \right)^2}{2} + \frac{a_1 \left(a_1 - 1 \right)}{2} \left(\epsilon_1' - \epsilon_1 \right)^2 \right. \\ &+ a_1 \left(\epsilon_0 \epsilon_1' - \epsilon_0 \epsilon_1 \right) \right] \end{split}$$

Squaring on both the sides and then taking expectation over the proposed estimator, we have

$$= S_y^4 (w_1 - 1)^2 + w_1^2 S_y^4 \left[E(\epsilon_0)^2 + a_1^2 E(\epsilon_1 - \epsilon_1')^2 + 2 a_1 E(\epsilon_0 \epsilon_1 - \epsilon_0 \epsilon_1') \right] + 2S_y^4 w_1 (w_1 - 1) \left[a_1 E(\epsilon_0 \epsilon_1' - \epsilon_0 \epsilon_1) - a_1 E(\epsilon_1 - \epsilon_1')^2 + \frac{a_1^2}{2} E(\epsilon_1 - \epsilon_1'))^2 \right]$$
(3.1)

Using the results from Sukhatme and Sukhatme, we have

 $E(\epsilon_0) = 0 = E(\epsilon_1), \ E(\epsilon_0)^2 = Ib_{2y}^*, \ E(\epsilon_1)^2 = Ib_{2x}^*, \ E(\epsilon_{1,1})^2 = I'b_{2x}^*, \ E(\epsilon_0\epsilon_1) = II_{22}^*, \ E(\epsilon_0\epsilon_1') = I'I_{22}^*, \ E($

Now, substituting the above results in (4.1), we get,

$$MSE(T_1^*) = S_y^4 w_1^2 \left[1 + Ib_{2y}^* + (I - I')(2a_1^2b_{2x}^* + 4a_1I_{22}^* - 2a_1b_{2x}^*) \right] - 2S_y^4 w_1 \left[1 + (I - I')(a_1I_{22}^* - a_1b_{2x}^* + \frac{a_1^2b_{2x}^*}{2}) \right] + S_y^4$$

Corollary 2. The optimum value of a_1 and w_1 are

$$a_1 = \frac{I_{22}^*}{b_{2x}}$$

$$w_1 = \frac{B_1}{A_1}$$

where

$$A_{1} = 1 + Ib_{2y}^{*} + (I - I')(2a_{1}^{2}b_{2x}^{*} + 4a_{1}I_{22}^{*} - 2a_{1}b_{2x}^{*})$$

$$B_{1} = 1 + (I - I')(a_{1}I_{22}^{*} - a_{1}b_{2x}^{*} + \frac{a_{1}^{2}b_{2x}^{*}}{2})$$

respectively. Also, the minimum mean squared error of T_1^* is

$$MSE(T_{1}^{*})_{opt} = S_{y}^{4}I\left[1-\frac{B_{1}^{2}}{A_{1}}\right]$$

4 Efficiency comparison

In this section, we compare the proposed classes of estimators with some important estimators. The comparison will be in terms of their MSEs up to the order of n^{-1} .

4.1 General estimator of population variance in case of SRSWOR

 $MSE(t_1) > MSE(T_1)_{opt}$

4.2 The ratio type variance estimator

 $MSE(t_2) > MSE(T_1)_{opt}$

4.3 Naik and Gupta (1996) ratio estimator using auxiliary attribute

 $MSE(t_3) > MSE(T_1)_{opt}$

4.4 The product estimator using auxiliary variables

 $MSE(t_4) > MSE(T_1)_{opt}$

Thus, it can be easily seen and verified that the proposed class of log-type estimators is far better than these above mentioned estimators, available in sampling literature.

5 Empirical study

The data on which we performed the numerical calculation is taken from some natural populations. The summary and the percent relative efficiency of the following estimators are as follows:

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Parameters	Population 1	Population 2	Population 3	Population 4		
n'	33	36	22	34		
n	11	11	8	10		
b_{2u}^{*}	4.032	2.632	9.433	2.725		
b_{2x}^*	1.388	2.402	7.105	12.366		
b_{2f}^{*}	1.143	2.345	7.766	1.912		
I_{22x}^*	0.305	1.835	8.140	0.224		
I_{22f}^{*}	1.155	2.014	8.538	2.104		
I_{22xf}^*	0.492	2.182	7.423	0.152		

Table 1: Parameters of the Data

Table 2: PRE of the estimators

Estimator	Population 1	Population 2	Population 3	Population 1
t_1	100	100	100	100
t_2	141.940	277.598	574168.1	637.142
t_3	38.525	13.593	13.009	12.407
t_4	116.860	248.212	5107.79	67.4228
T_{iopt}	142.235	293.078	605514.2	666.034

6 Conclusion

The present study extends the idea of Kumari et al. (2019) regarding the effective use of auxiliary information if the relationship between the study variable and the auxiliary variable is of logarithmic type. Further, the efficiency of the proposed estimators are compared with some conventional estimators and some recent estimators. The proposed estimator is most efficient than all the estimators. This study is also supported through an empirical sudy and the result of this study is quite encouraging.

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